

Chapter 3. Hitting times and Hitting Probabilities

Definitions

Definition 0.1 (First Hitting Time). Let $\mathbf{X} = (X_t)_{t \in \mathbb{N}}$ be a time-homogeneous Markov chain on a state space S . For any subset $A \subseteq S$, the **first hitting time** (or **first passage time**) to A is the random variable $\tau_A: \Omega \rightarrow \mathbb{N} \cup \{\infty\}$ defined as

$$\tau_A := \inf\{t \geq 0 : X_t \in A\}$$

where $\inf \emptyset = \infty$. If $A = \{x\}$ for a single state x , we often write τ_x instead of $\tau_{\{x\}}$.

Definition 0.2 (Hitting Probability). For $A \subset S$ and $x \in S$, the **hitting probability** of A from x is defined as

$$h_A(x) := \mathbb{P}_x(\tau_A < \infty)$$

This is the probability that the Markov chain, starting from state x , ever visits any state in the set A .

Main results

Theorem 0.1. The hitting probability $h_A(x) = \mathbb{P}_x(\tau_A < \infty)$ satisfies the following system of equations in the unknown h

$$\begin{cases} h(x) = 1 & \text{if } x \in A \\ ((I - P)h)(x) = 0 & \text{if } x \in A^c \end{cases} \quad (1)$$

Furthermore, $h_A(x)$ is the smallest non-negative solution to these equations. Namely if $g \geq 0$ solves the system Equation 1, then $g \geq h_A$.

Finally, h_A is the unique solution $g \geq 0$ to Equation 1 such that

$$\lim_{n \rightarrow \infty} \mathbb{E}_x[g(X_n) \mathbf{1}_{\tau_A > n}] = 0$$

Proof. We proceed step by step.

1. (Boundary condition) If $x \in A$, then $\tau_A = 0$ \mathbb{P}_x -a.s., and therefore $h_A(x) = 1$.
2. (Harmonicity) As $x \in A^c$, we have $\{\tau_A < \infty\} = \{\exists t \geq 0 : X_t \in A\} = \{\exists t \geq 1 : X_t \in A\}$ with \mathbb{P}_x -probability 1. Thus conditioning on the first step of the chain and using the [Markov property](#) for the function $\mathbf{1}_{\{\tau_A < \infty\}}$

$$\begin{aligned} h_A(x) &= \mathbb{P}_x(\tau_A < \infty) = \sum_{y \in S} \mathbb{P}_x(\tau_A < \infty \mid X_1 = y) \mathbb{P}_x(X_1 = y) \\ &= \sum_{y \in S} \mathbb{P}_y(\tau_A < \infty) p_{x,y} = \sum_{y \in S} p_{x,y} h_A(y) \end{aligned}$$

3. (Minimality) Let $g: S \rightarrow [0, \infty)$ be any non-negative function satisfying Equation 1. Let us prove by induction that

$$g(x) = \mathbb{P}_x(\tau_A \leq n) + \mathbb{E}_x[g(X_n)\mathbf{1}_{\tau_A > n}] \quad (2)$$

- Take $n = 1$. If $x \in A$, then the claim reduces to $1 = 1 + 0$. If $x \notin A$, then from Equation 1

$$g(x) = \mathbb{E}_x[g(X_1)] = \mathbb{E}_x[\mathbf{1}_A(X_1)] + \mathbb{E}_x[g(X_1)\mathbf{1}_{A^c}(X_1)] = \mathbb{E}_x[\mathbf{1}_{A^c}(X_0)\mathbf{1}_A(X_1)] + \mathbb{E}_x[g(X_1)\mathbf{1}_{A^c}(X_0)\mathbf{1}_{A^c}(X_1)]$$

which is equivalent to the claim by the definition of τ_A .

- Assume the claim for a given n . Since $\{\tau_A > n\} \subset \{X_n \notin A\}$, $g(X_n) = \mathbb{E}_{X_n}[g(X_1)]$, and reasoning as in the case $n = 1$

$$\begin{aligned} \mathbb{E}_x[g(X_n)\mathbf{1}_{\tau_A > n}] &= \mathbb{E}_x[\mathbb{E}_{X_n}[g(X_1)]\mathbf{1}_{\tau_A > n}] = \mathbb{E}_x[g(X_{n+1})\mathbf{1}_{\tau_A > n}] \\ &= \mathbb{E}_x[\mathbf{1}_{X_{n+1} \in A}\mathbf{1}_{\tau_A > n}] + \mathbb{E}_x[g(X_{n+1})\mathbf{1}_{X_{n+1} \notin A}\mathbf{1}_{\tau_A > n}] = \mathbb{E}_x[\mathbf{1}_{\tau_A = n+1}] + \mathbb{E}_x[g(X_{n+1})\mathbf{1}_{\tau_A > n+1}] \end{aligned}$$

From the inductive hypothesis and the last equality

$$g(x) = \mathbb{P}_x(\tau_A \leq n) + \mathbb{E}_x[\mathbf{1}_{\tau_A = n+1}] + \mathbb{E}_x[g(X_{n+1})\mathbf{1}_{\tau_A > n+1}] = \mathbb{P}_x(\tau_A \leq n+1) + \mathbb{E}_x[g(X_{n+1})\mathbf{1}_{\tau_A > n+1}]$$

Equation 2 is thus established. By monotone convergence, $\lim_n \mathbb{P}_x(\tau_A \leq n) = \mathbb{P}_x(\tau_A < \infty) = h_A(x)$. Therefore

$$g(x) = h_A(x) + \lim_n \mathbb{E}_x[g(X_n)\mathbf{1}_{\tau_A > n}] \geq h_A(x)$$

4. (Uniqueness) From the last equation, we see that the limit in that equation vanishes iff $g = h_A$. □

Examples

Example 0.1.

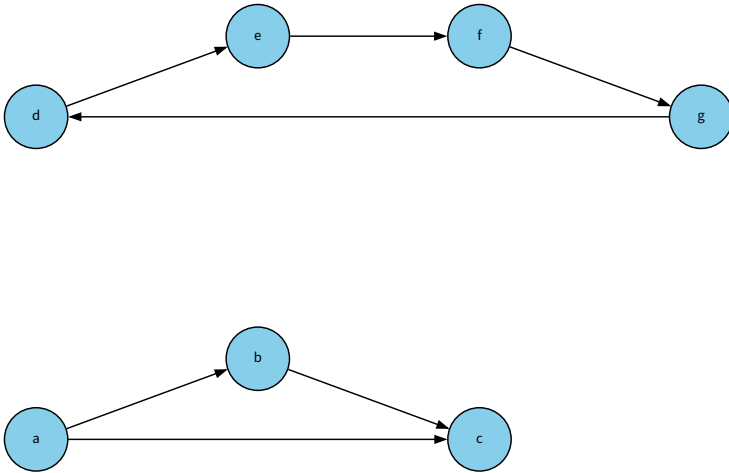


Figure 1

In this graph, the arrows represent strictly positive transition probabilities. Consider $h_b(x)$ the probability of hitting b when starting in x . On $\{a, b, c\}$ the function $h_b(x)$ is uniquely determined by the linear system Equation 1. On the other hand, on $\{d, e, f, g\}$ (which is disconnected from b), $h_b(x) = 0$; but $h(x) = c$ for $x \in \{d, e, f, g\}$ solves the harmonic equation Equation 1 for any constant c . We see indeed that h_b is the minimal non-negative solution to the system Equation 1.

Exercise 0.1. Let $A, B \subset S$ with $A \cap B = \emptyset$. Find the equation similar to Equation 1 satisfied by

$$h_{A,B}(x) := \mathbb{P}_x(\tau_A < \tau_B, \tau_{A \cup B} < \infty)$$

using two different methods

- Conditioning on the first step, as in the proof of Theorem 0.1.
- Considering a modified Markov chain with transition probabilities

$$q_{x,y} := \begin{cases} p_{x,y} & \text{if } x \notin B \\ \mathbf{1}_x(y) & \text{if } x \in B \end{cases}$$

Exercise 0.2. We want to give an analytical characterization of the transition probabilities $q_{x,y}$ of the trace of a recurrent Markov chain. We proved that the trace of \mathbf{X} on A has transition probabilities, for $x, y \in A$

$$q_{x,y} = \mathbb{P}_x(X_{\tau_A^+} = y)$$

Define the operator $P^{(A^c)}$ with entries

$$p_{x,y}^{(A^c)} = \begin{cases} p_{x,y} & \text{if } y \in A^c \\ 0 & \text{if } y \in A \end{cases}$$

Prove that, for each fixed $y \in A$, $q_{x,y}$ is the minimal solution to the problem in the unknown $h \geq 0$

$$((I - P^{(A^c)})h)(x) = p_{x,y}, \quad x \in S \tag{3}$$

meaning $q_{x,y} - \sum_{z \in A^c} p_{x,z} q_{z,y} = p_{x,y}$ for all $x \in S$.

Solution

Fix $y \in A$, and denote $h(x) \equiv q_{x,y} = \mathbb{P}_x(X_{\tau_A^+} = y)$ for $x \in S$. By conditioning on the first step:

$$h(x) = \sum_{z \in S} p_{x,z} \mathbb{P}_x(X_{\tau_A^+} = y | X_1 = z) = \sum_{z \in A} p_{x,z} \mathbf{1}_{y=z} + \sum_{z \in A^c} p_{x,z} h(z) = p_{x,y} + \sum_{z \in A^c} p_{x,z} h(z)$$

To prove minimality, let now $g \geq 0$ be any non-negative solution to the system Equation 3 for a fixed $y \in A$ (which we omit from the notation). Define $h^{(n)}(x) := \mathbb{P}_x(X_{\tau_A^+} = y, \tau_A^+ \leq n)$. Then $h^{(n)}(x) \leq q_{x,y}$, and by the continuity of probabilities on monotone sequences, $h^{(n)}(x) \uparrow h(x)$. Therefore it is enough to check that $g \geq h^{(n)}$ for all $n \geq 0$. Now we proceed by induction.

- Trivially $g \geq h^{(0)} \equiv 0$.
- $h^{(n)}$ satisfies (reasoning as for $q_{\cdot,y}$ above)

$$h^{(n+1)}(x) = p_{x,y} + \sum_{z \in A^c} p_{x,z} h^{(n)}(z)$$

Therefore by the induction hypothesis

$$g(x) = p_{x,y} + (P^{(A^c)}g)(x) \geq p_{x,y} + (P^{(A^c)}h^{(n)})(x) = h^{(n+1)}(x)$$

Exercise 0.3 (Gambler's ruin). A person goes to the casino with an initial capital of x dollars. At each game, they win a dollar with probability p and lose a dollar with probability $1 - p$. Their strategy is not to leave until they have a capital of N dollars, or they are broke. Find the probability that the person will be ruined in the game.

Exercise 0.4 (Birth and Death). We observe a population of bacteria and note each time one of them reproduces (the total population increases by 1) or dies (the total population decreases by 1). The probability of increase/decrease of the population depends on the population size. Say that if the population at a given moment is $n \geq 1$, the probability that a reproduction happens before a death is r_n (and a death happens before a reproduction with probability $1 - r_n$). Find the probability that the bacteria will go extinct, as a function of the sequence $(r_n)_{n \geq 1}$.

Exercise 0.5. You are walking in the desert, without water and with 2000 dollars. You find a magic lamp, and as you take it an evil magic Italian merchant appears, selling bottled water. You think you are safe, but there is a catch. One bottle costs 10000 dollars.

The Italian thus offers you the following game: you can bet any amount of money b on a six-face die roll. If you get 5 or 6 (thus with probability $p = 1/3$), you win b . Otherwise your money decreases by b . You can play as many rounds as you want, until you run out of money or decide to stop. Your goal is to maximize the probability of buying a bottle of water for the *cheap* price offered. Compute such an optimal probability.

i Abstraction

Hitting probabilities

In the context of generic Markov chains, we can define a stopping time $\tau: \Omega \rightarrow \Theta$ as a map such that $\{\tau \leq t\}$ is \mathcal{F}_t -measurable for all $t \in \Theta$. One can then define the σ -algebra of the events measurable up to the time τ :

$$\mathcal{F}_\tau := \{A \in \mathcal{F} : A \cap \{\tau \leq t\} \in \mathcal{F}_t \text{ for all } t \in \Theta\}$$

If we take $\Theta = \mathbb{N} \cup \{\infty\}$ or $\Theta = [0, \infty]$, one says that the strong Markov Property holds if

$$\mathbb{E}_\mu[F(\theta_\tau \mathbf{X}) \mid X_\tau = x, \tau < \infty] = \mathbb{E}_x[F(\mathbf{X})]$$

for all bounded measurable maps $F: D(E) \rightarrow \mathbb{R}$ and for all stopping times such that $\mathbb{P}_\mu(\tau < \infty) > 0$. Notice that in this general framework, one needs some further assumptions to guarantee that this holds, it is not just a consequence of the Markov property. From which the name **strong** Markov. In this sense, **we proved** that a discrete-time ($\Theta = \mathbb{N}$) Markov chain on a countable state space (or indeed on Polish spaces with minor adaptations) has the strong Markov property.